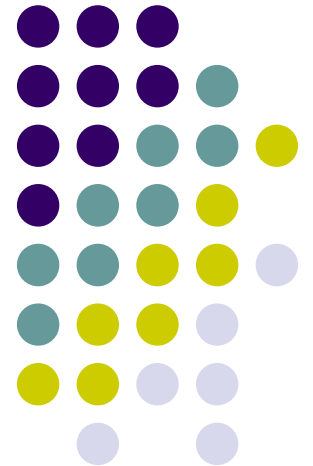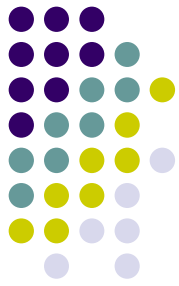# Hidden Markov Models

Zhang Hongxin

zhx@cad.zju.edu.cn

State Key Lab of CAD&CG, ZJU

2006-04-13

Outline

- Background

- Markov Chains

- Hidden Markov Models

Example: Video Texture

- Problem statement

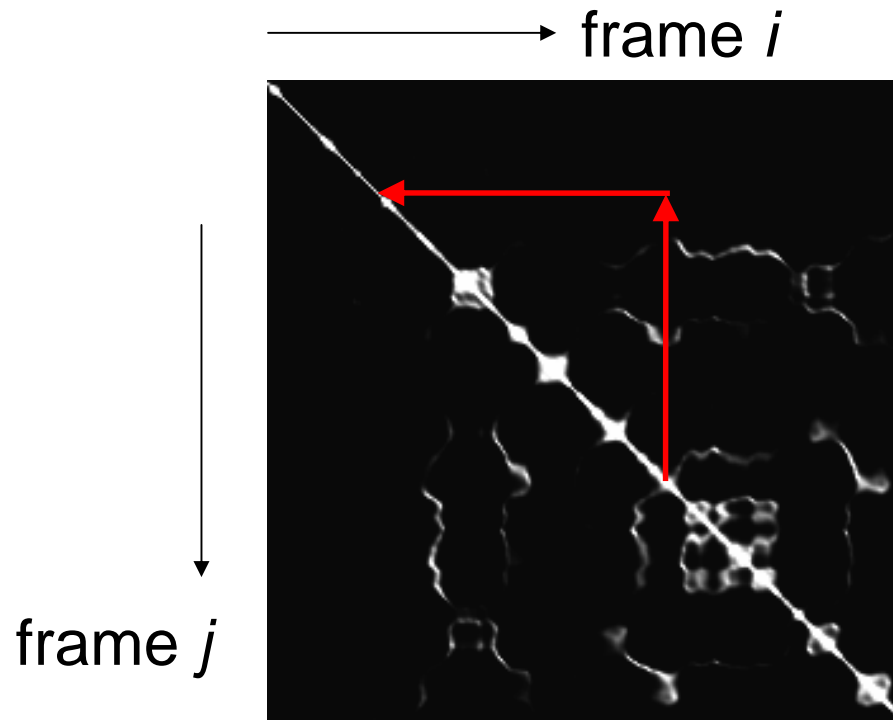video clip   video texture

The approach

How do we find good transitions?

Finding good transitions

Compute L_2 distance $D_{i,\,j}$ between all frames

frame i

frame j

Similar frames make good transitions

Demo: Fish Tank

Mathematic model of Video Texture

A sequence of random variables

{ADEABEDADBCAD}

A sequence of random variables

{BDACBDCACDBCADCBADCA}

Markov Model

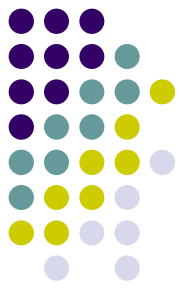The future is independent of the past and given by the present.

Markov Property

- Formal definition

  - Let $X=\{X_n\}_{n=0\ldots N}$ be a sequence of random variables taking values $s_k \in N$ if and only if
    $$P(X_m=s_m/X_0=s_0,\ldots,X_{m-1}=s_{m-1}) = P(X_m=s_m/ X_{m-1}=s_{m-1})$$

  then the X fulfills Markov property

- Informal definition

  - The future is independent of the past given the present.

History of MC

- Markov chain theory developed around 1900.
- Hidden Markov Models developed in late 1960's.
- Used extensively in speech recognition in 1960-70.
- Introduced to computer science in 1989.

Andrei Andreyevich Markov

Applications

- ➢ Bioinformatics.
- ➢ Signal Processing
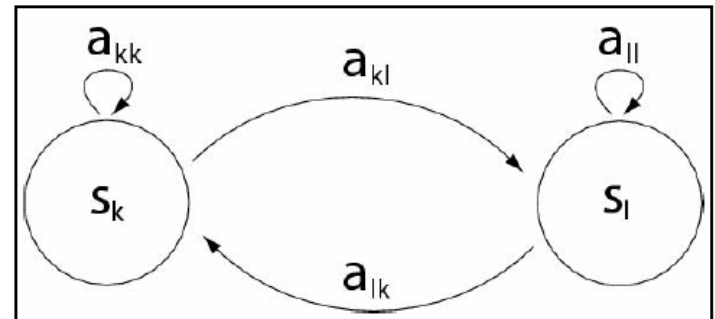- ➢ Data analysis and Pattern recognition

Markov Chain

- A Markov chain is specified by
  - A state space $S = \{ s_1, s_2..., s_n \}$
  - An initial distribution a_0
  - A transition matrix A

  Where $A(n)_{ij} = a_{ij} = P(q_t=s_j/q_{t-1}=s_i)$

- Graphical Representation

  as a directed graph where
  - Vertices represent states
  - Edges represent transitions with positive probability

Probability Axioms

- Marginal Probability – sum the joint probability

$$P(x = a_i) \equiv \sum_{y \in A_Y} P(x = a_i, y)$$

- Conditional Probability

$$P(x = a_i \mid y = b_j) \equiv \frac{P(x = a_i, y = b_j)}{P(y = b_j)} \text{ if } P(y = b_j) \neq 0.$$

Calculating with Markov chains

- Probability of an observation sequence:
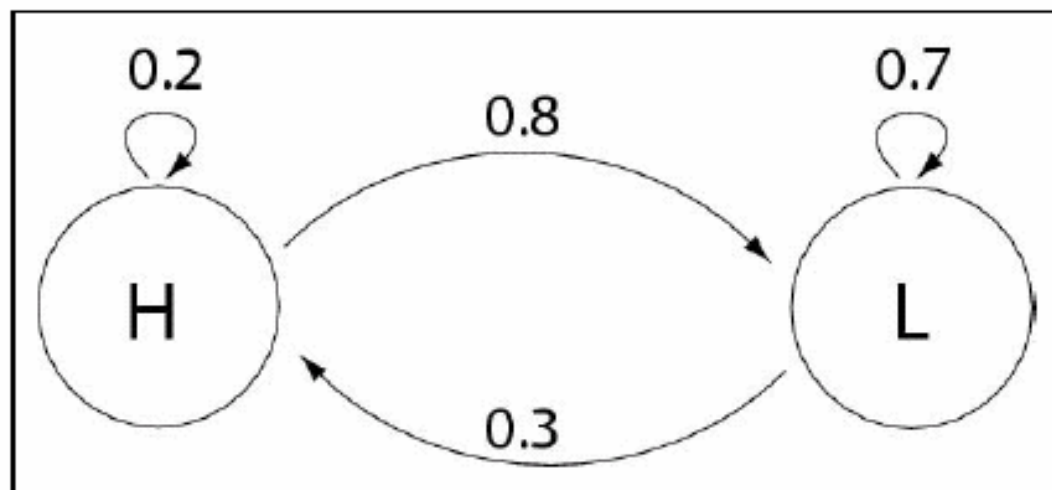 - Let $X=\{x_t\}^L_{t=0}$ be an observation sequence from the Markov chain $\{S, a_0, A\}$

$$P(x) = P(x_L, \ldots, x_1, x_0)$$
$$= P(x_L \mid x_{L-1}, \ldots, x_0)P(x_{L-1} \mid x_{L-2}, \ldots, x_0) \cdots P(x_0)$$
$$= P(x_L \mid x_{L-1})P(x_{L-1} \mid x_{L-2}) \cdots P(x_0)$$
$$= \mathbf{b}_{x_0} \prod_{i=1}^{L} a_{x_{i-1}x_i}$$

Example

Assume we are modeling a time series of high and low pressures during the Danish autumn.

Let $S = \{H, L\}$, $\mathbf{b} = \pi = \begin{bmatrix} \dfrac{3}{11}, \dfrac{8}{11} \end{bmatrix}$, and $A = \begin{bmatrix} 0.2 & 0.8 \\ 0.3 & 0.7 \end{bmatrix}$.
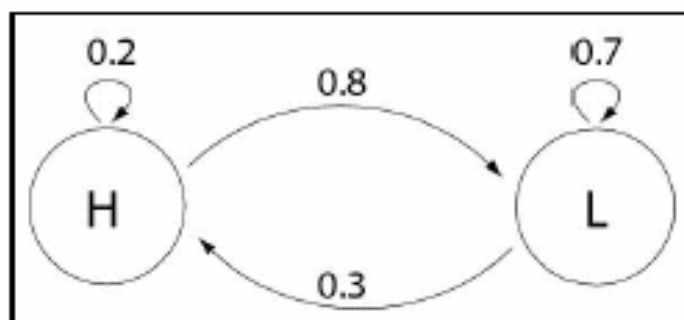
Graphical representation of A

Example

Comparing likelihoods

We want to know the likelihood of one week of high pressure in Denmark (DK) versus California (Cal).

x=HHHHHHH

$$P(x \mid DK)$$

$$= \mathbf{b}_H a_{HH} a_{HH} a_{HH} a_{HH} a_{HH} a_{HH} a_{HH}$$

$$= \frac{3}{11} \left(\frac{1}{5} \right)^6 \approx 0.0017\%$$

Motivation of Hidden Markov Models

- Hidden states
  - The state of the entity we want to model is often not observable:
  - The state is then said to be hidden.

- Observables
  - Sometimes we can instead observe the state of entities influenced by the hidden state.

- A system can be modeled by an HMM if:
  - The sequence of hidden states is Markov
  - The sequence of observations are independent (or Markov) given the hidden

Hidden Markov Model

- Definition $M = \{S, V, A, B, \pi\}$
  - **Set of states** $\qquad S = \{s_1, s_2, \ldots, s_N\}$
  - **Observation symbols** $\quad V = \{v_1, v_2, \ldots, v_M\}$
  - **Transition probabilities**
    - A between any two states $\quad a_{ij} = P(q_t = s_j | q_{t-1} = s_i)$
  - **Emission probabilities**
    - B within each state $\quad b_j(O_t) = P(O_t = v_j | q_t = s_j)$
  - **Start probabilities** $\qquad \pi = \{a_0\}$

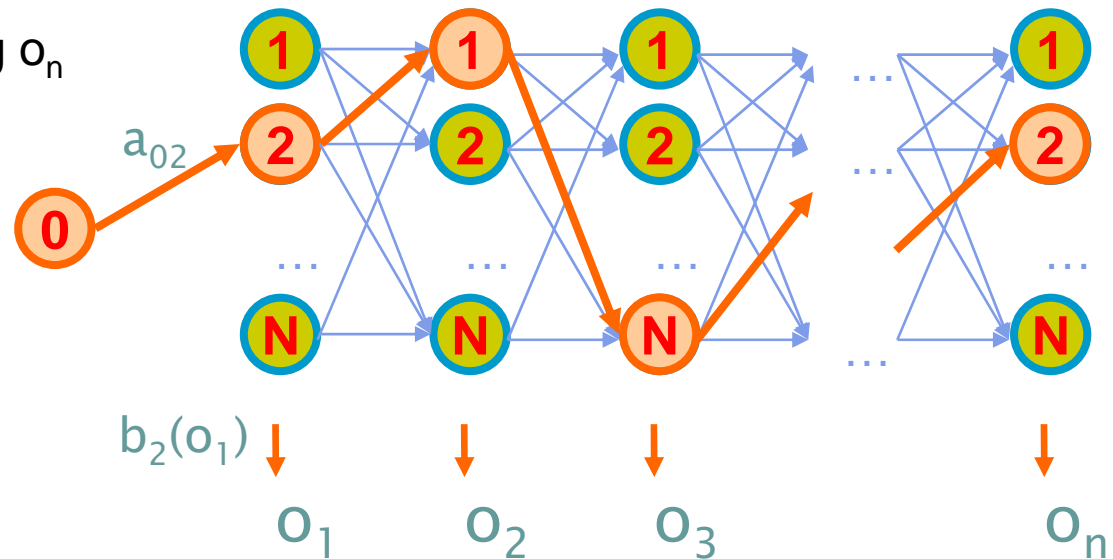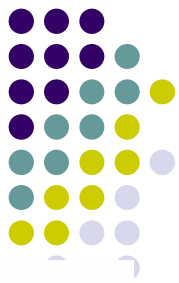Use $\lambda = (A, B, \pi)$ to indicate the parameter set of the model.
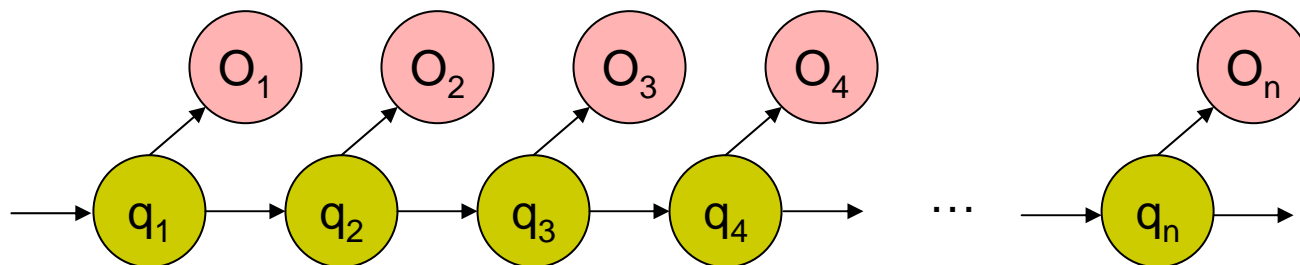
Generating a sequence by the model

Given a HMM, we can generate a sequence of length n as follows:

1. Start at state q_1 according to prob a_{0t_1}

2. Emit letter o_1 according to prob $e_{t_1}(o_1)$

3. Go to state q_2 according to prob $a_{t_1t_2}$
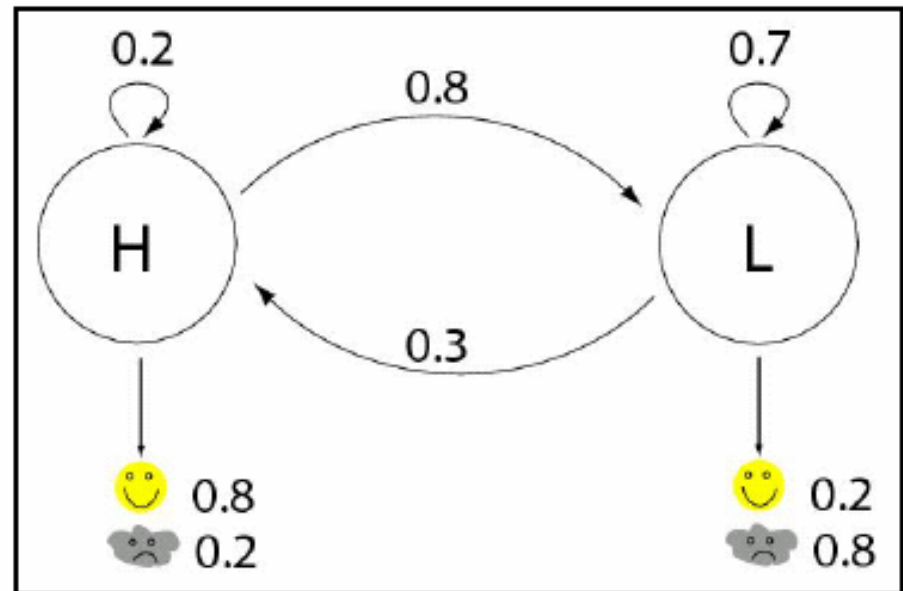
4. … until emitting o_n

Example

Model of high and low pressures

Assume we can not measure high and low pressures.

The state of the weather is influenced by the air pressure.

We make an HMM with hidden states representing high and low pressure and observations representing the weather:

Hidden states: L L L L H H L
Observations: ☹ ☹ ☺ ☹ ☺ ☺ ☹

Calculating with Hidden Markov Model
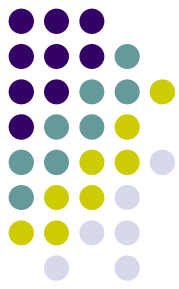
Consider one such fixed state sequence

$$Q = q_1 q_2 \cdots q_T$$

The observation sequence O for the Q is

$$P(O \mid Q, \lambda) = \prod_{t=1}^{T} P(O_t \mid q_t, \lambda)$$

$$= b_{q_1}(O_1) \cdot b_{q_2}(O_2) \cdots b_{q_T}(O_T)$$

Calculating with Hidden Markov Model (cont.)

The probability of such a state sequence Q can be written as

$$P(Q \mid \lambda) = a_{0q_1} a_{q_1 q_2} \cdot a_{q_2 q_3} \cdots a_{q_{T-1} q_T}$$

$$P(O, Q \mid \lambda) = P(O \mid Q, \lambda) P(Q \mid \lambda)$$

The probability that O and Q occur simultaneously,
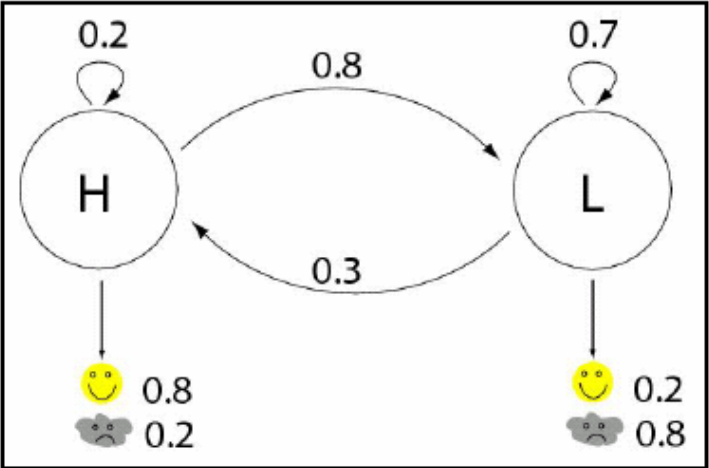is simply the product of the above two terms, i.e.,

$$P(O, Q \mid \lambda) = a_{0q_1} b_{q_1}(O_1) a_{q_1 q_2} b_{q_2}(O_2) a_{q_2 q_3} \cdots a_{q_{T-1} q_T} b_{q_T}(O_T)$$

Example

$$P(x,\pi)$$
$$= \left(a_{0L}e_L(R)\right)\left(a_{LL}e_L(R)\right)\left(a_{LL}e_L(S)\right)\left(a_{LL}e_L(R)\right)\left(a_{LH}e_H(S)\right)\left(a_{HH}e_H(S)\right)\left(a_{HL}e_L(R)\right)$$
$$= \left(\frac{8}{11}\frac{8}{10}\right)\left(\frac{7}{10}\frac{8}{10}\right)\left(\frac{7}{10}\frac{2}{10}\right)\left(\frac{7}{10}\frac{8}{10}\right)\left(\frac{3}{10}\frac{8}{10}\right)\left(\frac{2}{10}\frac{8}{10}\right)\left(\frac{8}{10}\frac{8}{10}\right)$$
$$= 0.0006278$$

Hidden states: L L L L H H L
Observations:

The three main questions on HMMs

1. **Evaluation**
   GIVEN   a HMM $M=(S, V, A, B, \pi)$, and a sequence O,
   FIND $P[O|M]$

2. **Decoding**
   GIVEN   a HMM $M=(S, V, A, B, \pi)$, and a sequence O,
   FIND   the sequence Q of states that maximizes $P(O, Q \mid \lambda)$

3. **Learning**
   GIVEN   a HMM $M=(S, V, A, B, \pi)$, with unspecified
                transition/emission probabilities and a sequence Q,
   FIND   parameters $\theta = (e_i(.), a_{ij})$ that maximize $P[x|\theta]$

Evaluation

➢ Find the likelihood a sequence is generated by the model

➢ A straightforward way （穷举法）

    ➢ The probability of *O* is obtained by summing all possible state sequences *q* giving

Complexity is *O(N^T)*

Complexity is $O(N^T)$

Calculations is unfeasible

$$P(O \mid \lambda) = \sum_{all\ Q} P(O \mid Q, \lambda) P(Q \mid \lambda)$$

$$= \sum_{q_1, q_2, \ldots q_T} \pi_{q_1} b_{q_1}(O_1) a_{q_1 q_2} b_{q_2}(O_2) a_{q_2 q_3} \cdots a_{q_{T-1} q_T} b_{q_T}(O_T)$$

The Forward Algorithm
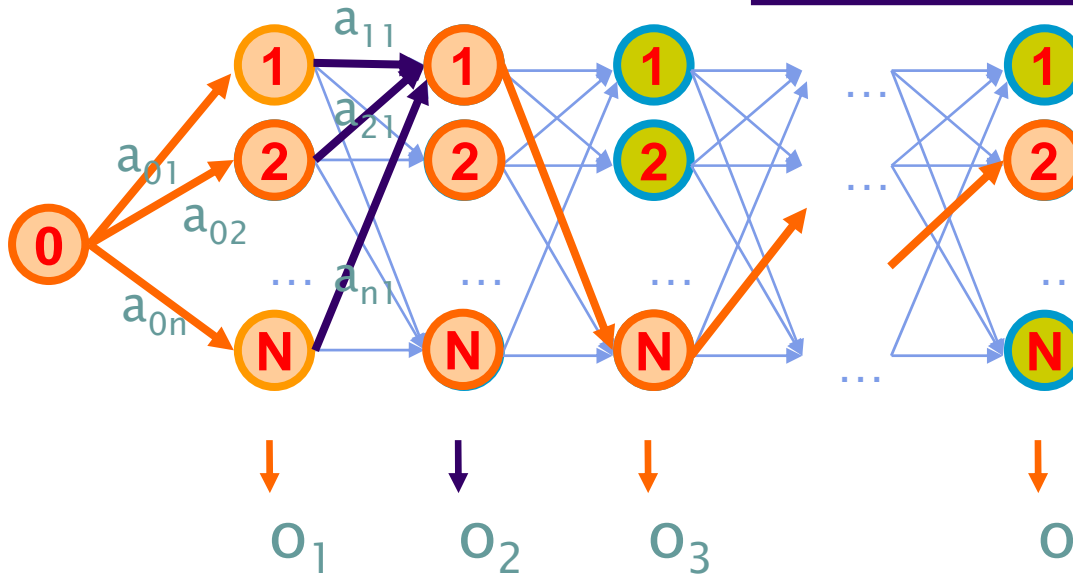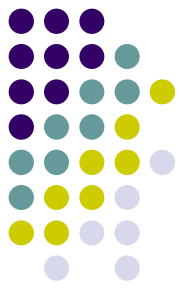
- A more elaborate algorithm
  - The Forward Algorithm

$$\alpha_2(1) = [\sum_{i=1}^{N} \alpha_1(i) a_{i1}] b_1(O_2)$$

$$P(O_1 O_2 \mid \lambda) = \sum_{i=1}^{N} \alpha_2(i)$$

$$P(O \mid \lambda) = \sum_{i=1}^{N} \alpha_T(i)$$

The Forward Algorithm

The Forward variable

$$\alpha_t(i) = P(O_1 O_2 \cdots O_t, q_t = S_i \mid \lambda)$$

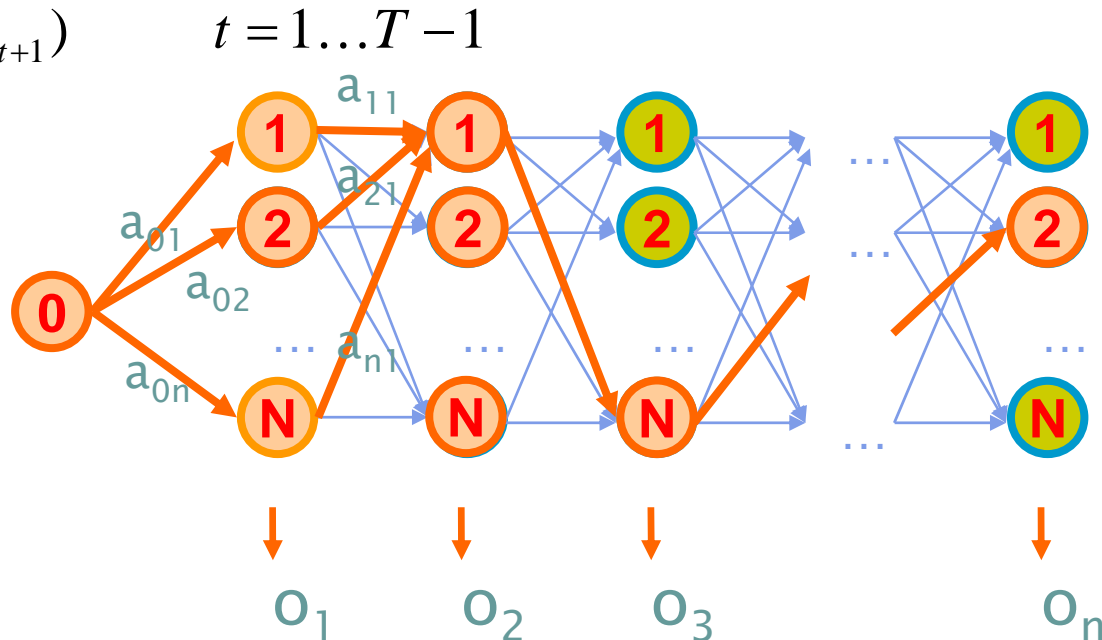We can compute $\alpha(i)$ for all N, i,

Initialization:

$$\alpha_1(i) = a_i b_i(O_1) \quad i = 1 \ldots N$$

Iteration:

$$\alpha_{t+1}(i) = [\sum_{i=1}^{N} \alpha_t(i) a_{ij}] b_j(O_{t+1}) \quad t = 1 \ldots T - 1$$

Termination:

$$P(O \mid \lambda) = \sum_{i=1}^{N} \alpha_T(i)$$

The Backward Algorithm

The backward variable

$$\beta_t(i) = P(O_{t+1}O_{t+2}\cdots O_T \mid q_t = S_i, \lambda)$$

Similar, we can compute backward variable for all *N, i,*

Initialization:

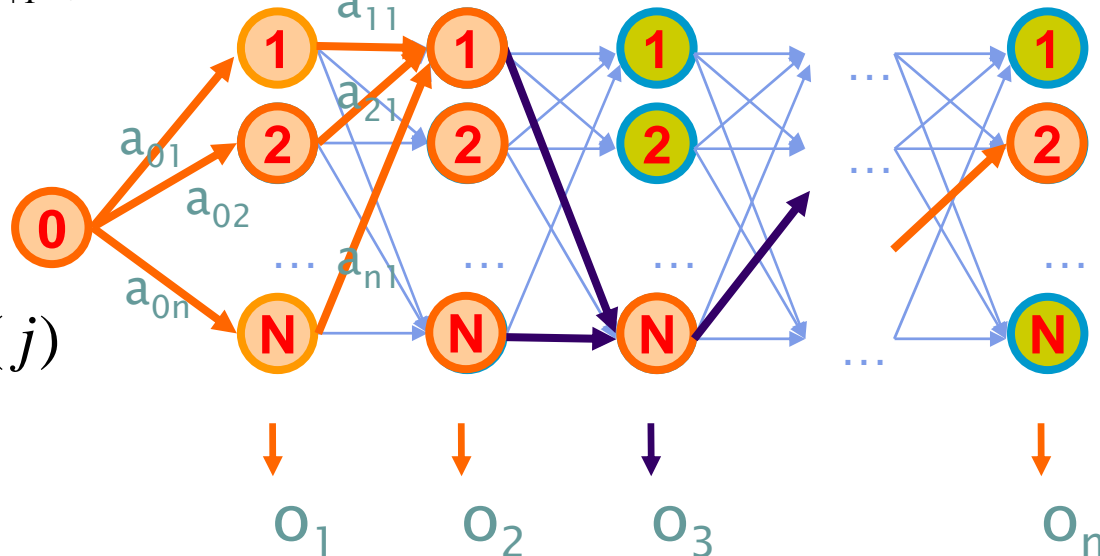$$\beta_T(i) = 1, \ i = 1,...,N$$

Iteration:

$$\beta_t(i) = \sum_{j=1}^{N} a_{ij} b_j(O_{t+1}) \beta_{t+1}(j) \qquad t = T-1, T-2, \cdots, 1, 1 \le i \le N$$

Termination:

$$P(O \mid \lambda) = \sum_{j=1}^{N} a_{0j} b_1(O_1) \beta_1(j)$$

Consider $\alpha_T(i) = P(O_1 O_2 \ldots O_T, q_T = S_i \mid \lambda)$

Thus $P(q_T = S_i \mid O) = \dfrac{P(O, q_T = S_i)}{P(O)} = \dfrac{\alpha_T(i_T)}{\displaystyle\sum_i \alpha_T(i_T)}$

Also $P(q_t = S_i \mid O) = \dfrac{P(O, q_t = S_i)}{P(O)}$

$$= \dfrac{P(O_1 O_2 \cdots O_t, q_t = S_{i_t}, O_{t+1} O_{t+2} \cdots O_T)}{P(O)}$$

$$= \dfrac{P(O_1 O_2 \cdots O_t, q_t = S_i) P(O_{t+1} O_{t+2} \cdots O_T \mid O_1 O_2 \cdots O_t, q_t = S_i)}{P(O)}$$

Forward, f$_k$(i)   **Backward, b$_k$(i)**

$$= \dfrac{P(O_1 O_2 \cdots O_t, q_t = S_i) \, P(O_{t+1} O_{t+2} \cdots O_T \mid q_t = S_i)}{P(O)}$$

$$= \dfrac{\alpha_t(i) \beta_t(i)}{\displaystyle\sum_i \alpha_T(i)} = \gamma(i)$$

Decoding

GIVEN   a HMM, and a sequence O.

Suppose that we know the parameters of the Hidden Markov Model and the observed sequence of observations $O_1, O_2, \ldots, O_T$.

FIND   the sequence Q of states that maximizes $P(Q/O, \lambda)$

Determining the sequence of States $q_1, q_2, \ldots, q_T$, which is optimal in some meaningful sense. (i.e. best "explain" the observations)
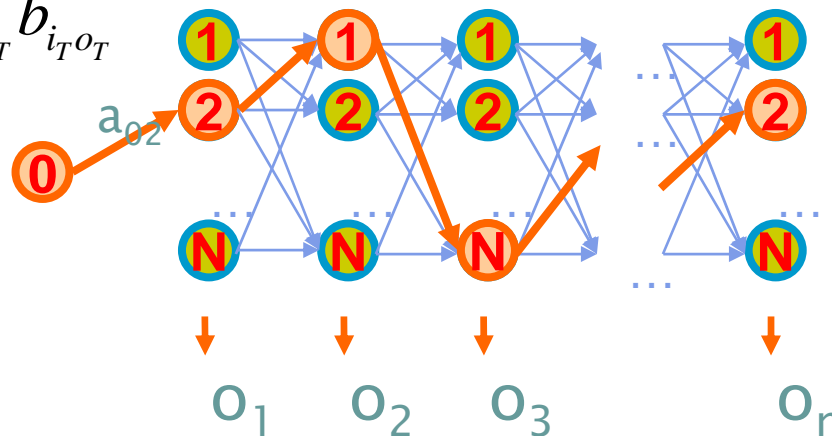
Decoding

Consider $P(Q|O, \lambda) = \dfrac{P(O, Q \mid \lambda)}{P(O \mid \lambda)}$

To maximize the above probability is equivalent to maximizing

$P(O, Q \mid \lambda)$

$= a_{i_1} b_{i_1 o_1} a_{i_1 i_2} b_{i_2 o_2} a_{i_2 i_3} b_{i_3 o_3} \dots a_{i_{T-1} i_T} b_{i_T o_T}$

A best path finding problem

$\max P(O, Q \mid \lambda)$

$= \max \ln(P(O, Q \mid \lambda))$

$= \max(\ln\!\left(a_{i_1} b_{i_1 o_1}\right) + \ln\!\left(a_{i_1 i_2} b_{i_2 o_2}\right) \dots + \ln\!\left(a_{i_{T-1} i_T} b_{i_T o_T}\right))$

Viterbi Algorithm

[Dynamic programming]

Initialization:

$\delta_1(i) = a_{0i}b_i(O_1)$, $i = 1 \ldots N$

$\psi_1(i) = 0.$

Recursion:

$\delta_t(j) = \max_i [\delta_{t-1}(i)\, a_{ij}]b_j(O_t)$ $t=2 \ldots T$ $j=1 \ldots N$

$\psi_1(j) = \text{argmax}_i [\delta_{t-1}(i)\, a_{ij}]$ $t=2 \ldots T$ $j=1 \ldots N$

Termination:

$P^* = \max_i \delta_T(i)$

$q_T^* = \text{argmax}_i [\delta_T(i)]$

Traceback:

$q_t^* = \psi_1(q^*_{t+1})$ $t=T\text{-}1, T\text{-}2, \ldots, 1.$
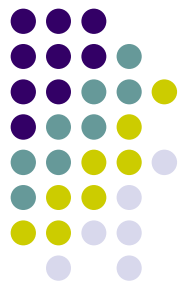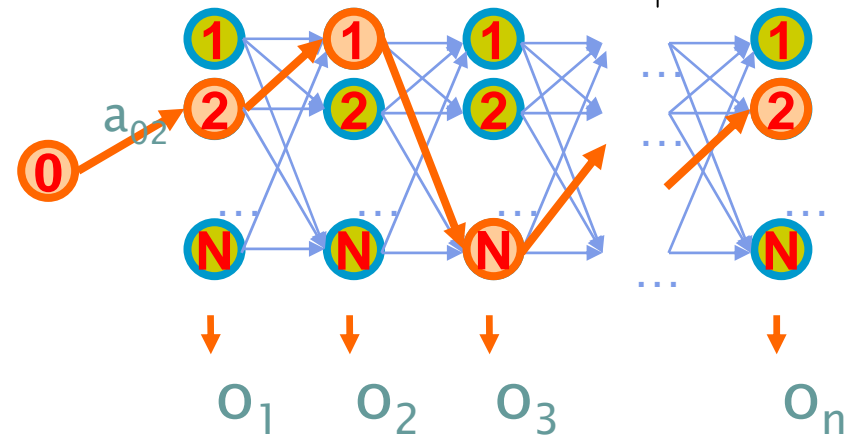
The Viterbi Algorithm

x_1 x_2 x_3 ....x_N
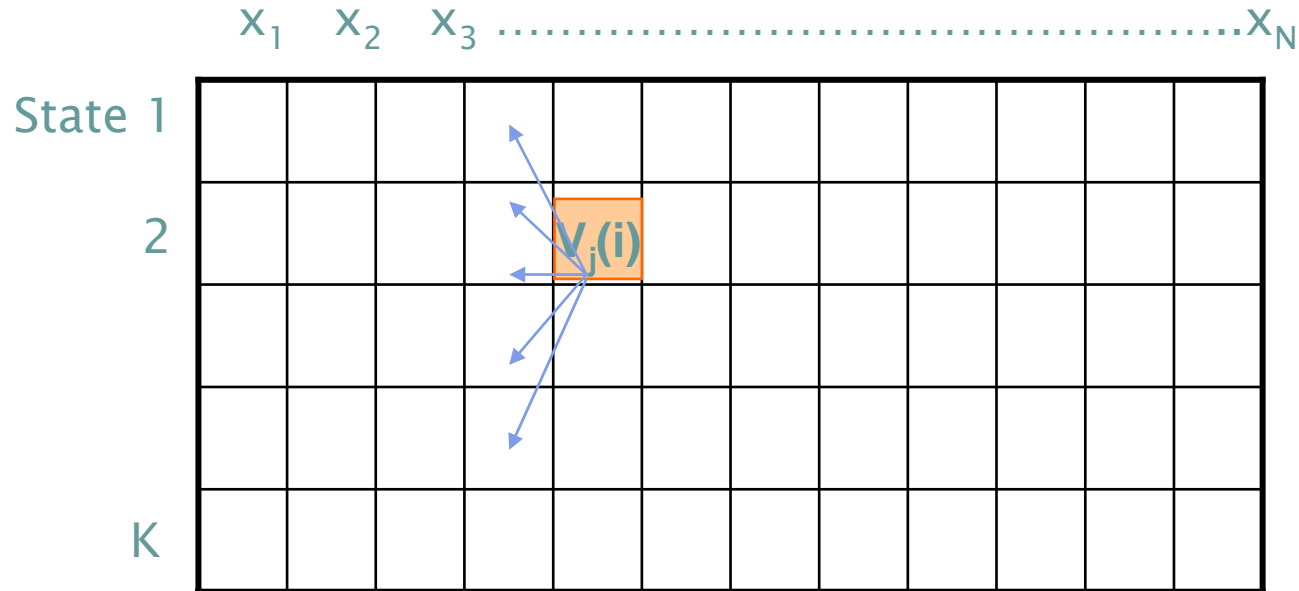
Similar to "aligning" a set of states to a sequence

Time: $O(K^2N)$

Space: $O(KN)$

Learning

- Estimation of Parameters of a Hidden Markov Model

    1. Both the sequence of observations O and the sequence of States Q is observed

       learning $\lambda = (A, B, \pi)$

    2. Only the sequence of observations O are observed

       learning Q and $\lambda = (A, B, \pi)$

Maximal Likelihood Estimation

- Given O and Q, the Likelihood is given by:

$$L(A, B, \pi) = a_{i_1} b_{i_1 o_1} a_{i_1 i_2} b_{i_2 o_2} a_{i_2 i_3} b_{i_3 o_3} \ldots a_{i_{T-1} i_T} b_{i_T o_T}$$

- the log-Likelihood is:

$$l(A, B, \pi) = \ln L(A, B, \pi) = \ln(a_{i_1}) + \ln(b_{i_1 o_1}) + \ln(a_{i_1 i_2})$$

$$+ \ln(a_{i_2 i_3}) + \ln(b_{i_3 o_3}) \ldots + \ln(a_{i_{T-1} i_T}) + \ln(b_{i_T o_T})$$

$$= \sum_{i=1}^{M} f_{i0} \ln(a_i) + \sum_{i=1}^{M} \sum_{j=1}^{M} f_{ij} \ln(a_{ij}) + \sum_{i=1}^{M} \sum_{o(i)} \ln(b_{io})$$

where $f_{i0} = $ the number of times state i occurs in the first state

$f_{ij} = $ the number of times state i changes to state j.

$\beta_{iy} = f(y|\theta_i)$ (or $p(y|\theta_i)$ in the discrete case)

$\sum_{o(i)} \bullet = $ the sum of all observations o_t where $q_t = S_i$

Maximal Likelihood Estimation

In such case these parameters computed by *Maximum Likelihood estimation are*:

$$\hat{a}_i = \frac{f_{i0}}{1} \qquad \hat{a}_{ij} = \frac{f_{ij}}{\sum_{j=1}^{M} f_{ij}}, \text{and}$$

$\hat{b}_i =$ the MLE of b_i computed from the observations o_t where $q_t = S_i$.

Maximal Likelihood Estimation

- Only the sequence of observations O are observed

$$L(A, B, \pi) = \sum_{i_1, i_2 \ldots i_T} a_{i_1} b_{i_1 o_1} a_{i_1 i_2} b_{i_2 o_2} a_{i_2 i_3} b_{i_3 o_3} \ldots a_{i_{T-1} i_T} b_{i_T o_T}$$

- It is difficult to find the Maximum Likelihood Estimates directly from the Likelihood function.

- The Techniques that are used are

  1. *The Segmental K-means Algorith*
  2. *The Baum-Welch (E-M) Algorithm*

The Baum-Welch Algorithm

- The E-M algorithm was designed originally to handle "Missing observations".

- In this case the missing observations are the states $\{q_1, q_2, \dots, q_T\}$.

- Assuming a model, the states are estimated by finding their expected values under this model. (The E part of the E-M algorithm).

The Baum-Welch Algorithm

- With these values the model is estimated by Maximum Likelihood Estimation (The M part of the E-M algorithm).

- The process is repeated until the estimated model converges.

The Baum-Welch Algorithm

Initialization:

Pick the best-guess for model parameters (or arbitrary)

Iteration:

Forward

Backward

Calculate A_{kl}, $E_k(b)$

Calculate new model parameters a_{kl}, $e_k(b)$

Calculate new log-likelihood $P(x \mid \theta)$

GUARANTEED TO BE HIGHER BY EXPECTATION-MAXIMIZATION

Until $P(x \mid \theta)$ does not change much

The Baum-Welch Algorithm

Let $f(O,Q|\lambda) = L(O,Q,\lambda)$ denote the joint distribution of Q,O.

Consider the function:

$$Q(\lambda, \lambda') = E_{\mathbf{X}}\left(\ln L(O,Q,\lambda)\big|Q,\lambda'\right)$$
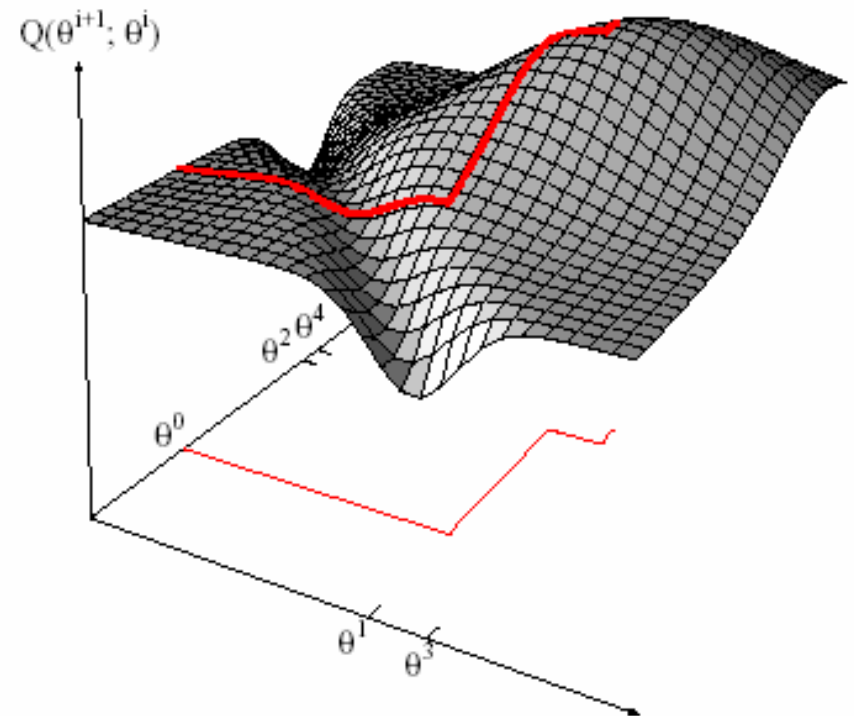
Starting with an initial estimate of λ $\left(\lambda^{(1)}\right)$. A sequence of estimates $\left\{\lambda^{(m)}\right\}$ are formed by finding $\lambda = \lambda^{(m+1)}$ to maximize $Q\left(\lambda, \lambda^{(m)}\right)$ with respect to λ.

The Baum-Welch Algorithm

The sequence of estimates $\left\{\lambda^{(m)}\right\}$ converge to a local maximum of the likelihood

$$L(Q, \lambda) = f(Q|\lambda)$$

Markov Random field

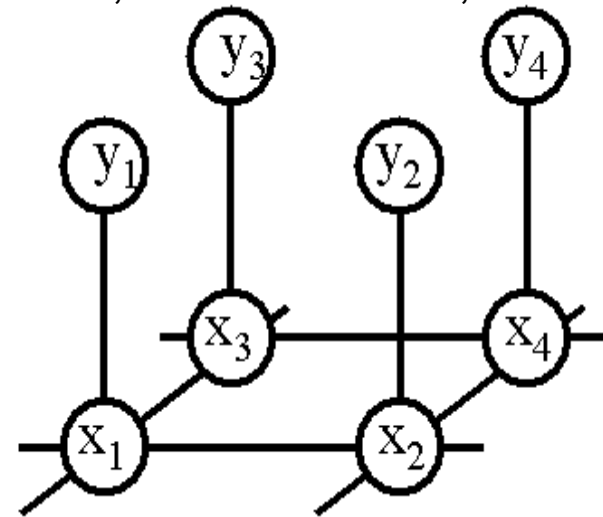- http://www.nlpr.ia.ac.cn/users/szli/MRF_Book/MRF_Book.html
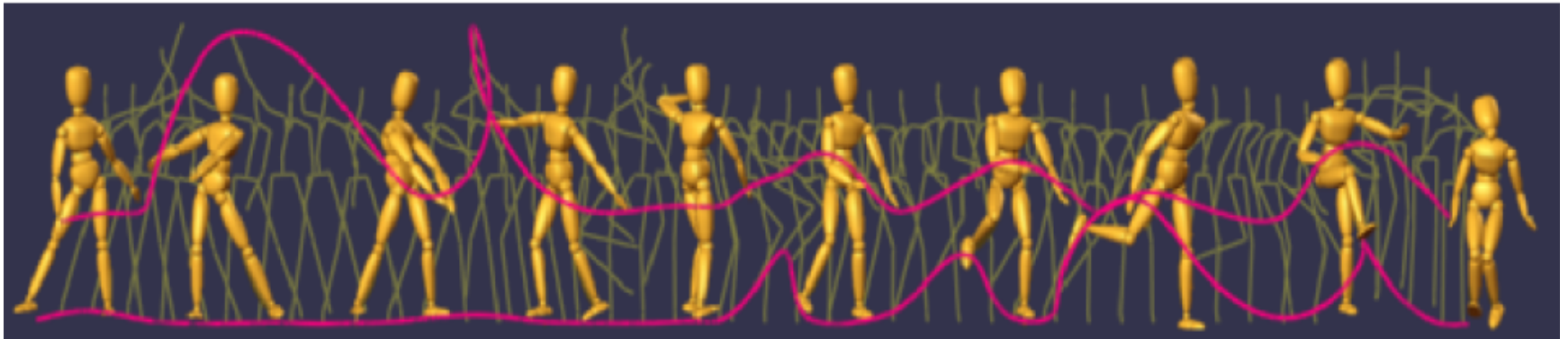
Belief Network (Propagation)

Y. Weiss and W. T. Freeman
Correctness of Belief Propagation in Gaussian Graphical Models of Arbitrary Topology. in: Advances in Neural Information Processing Systems 12, edited by S. A. Solla, T. K. Leen, and K-R Muller, 2000. MERL-TR99-38.

Motion Texture

- *Motion Texture: A Two-Level Statistical Model for Character Motion Synthesis.* Yan Li, Tianshu Wang, and Heung-Yeung Shum. SIGGRAPH 2002.
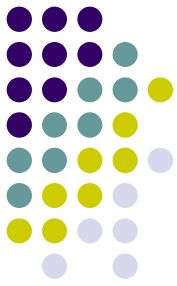
Plant Texture

(a)~(b): photographs. (c): recovered sample. (d)~(f): synthesized plants

Figure 8. Yulan

Homework

- Read the motion texture siggraph paper.